GUIDE FOR THE USE OF THE DECISION SUPPORT SYSTEM (DSS)*

El Adlouni, Salaheddine, Bernard Bobée and Ouejdene Samoud

bernard.bobee@ete.inrs.ca; el_adlouni@yahoo.com; ouejdene.samoud.1@ulaval.ca

1. Introduction to the DSS

Eighteen statistical distributions are available in HYFRAN-PLUS software to fit data sets that are independent, homogenous and stationary. A Decision Support System (DSS) is developed to help to the selection of the most appropriate class of distributions, with respect to extreme values. Distributions that are usually used in flood frequency analysis can be grouped in three categories that contain the ten distributions that are widely used in hydrology to represent maximum annual flow series:

- Class C (regularly varying distributions): Fréchet (EV2), Halphen IB (HIB), Log-Pearson (LP3), Inverse Gamma (IG).
- Class D (sub-exponential distributions): Halphen type A (HA), Halphen type B (HB), Gumbel (EV1), Pearson type III (PIII), Gamma (G).
- Class E (Exponential distribution).

Figure 1 presents exponential (E), sub-exponential (D) and regularly varying (C) distributions. Distributions are ordered from light tailed (from the left) to heavy tailed (to the right). The limiting cases (bottom squares) represented by distributions in the limits of classes. The tail of the class C distributions is heavier than that of the class D distributions, which is heavier than that of the class E. Thus, estimated quantiles can be ordered equivalently. Indeed, for a given sample, the T-event corresponds to the quantile of the probability of non-exceedance $p = 1 - 1/T$ estimated by distributions of the classes C, D and E, are QT (C), QT (D) and QT (E) respectively, which verify the following relation: QT (E) < QT (D) < QT (C).

The Lognormal distribution (LN) doesn't belong to any of these classes. It has an asymptotic behaviour which is in the frontier of the classes C and D. Indeed, the LN tail is lighter (respectively, heavier) than that of a distribution of the class C (respectively, class D). Thus, the quantiles (QT) estimated by a distribution belonging to the classes C, D and the LN, verify the following relation:

QT (D) < QT (LN) < QT (C) (Figure 1). Consequently:

- If the true distribution is regularly varying (class C), and the LN distribution is considered for the fit, there a risk to underestimate the quantiles;
- If the true distribution is sub-exponential (class D), and the LN distribution is considered for the fit, there a risk to overestimate the quantiles.

In the last version of the DSS, and to have a safe choice, LN is considered by default as a distribution of the class D. In the present version, a new step is added to the DSS in order to test Lognormality before the adequacy of the classes C or D.
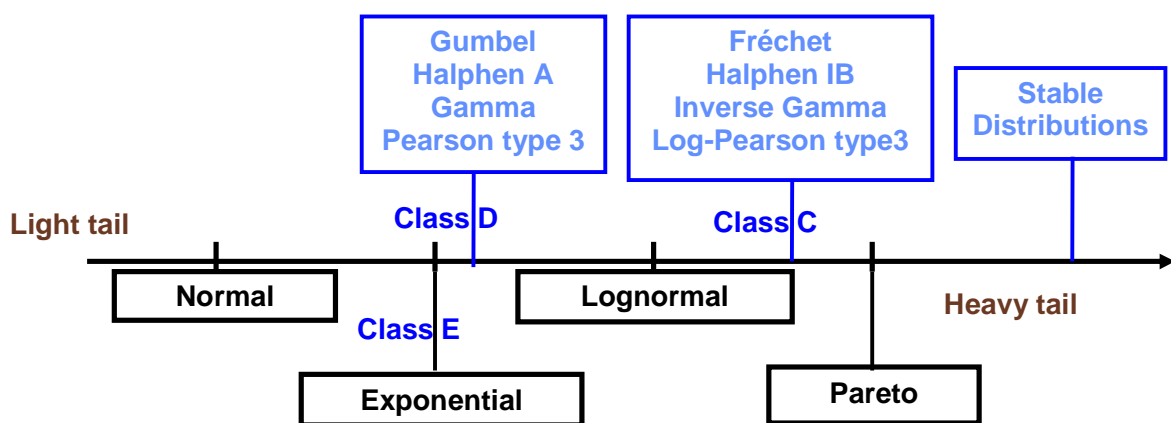

Figure 1: distributions ordered with respect to their right tails (El Adlouni et al., 2008).

The methods developed in the DSS allow the identification of the most adequate class of distribution to fit a given sample, especially for extremes. These methods are (cf. Diagram):

- Log-normality test : Use the (Cv,Cs) diagram and then the Jarque-Bera test if it is recommended;
- The Log-Log plot : used to discriminate between on the one hand the class C and on the other hand the classes E and D;
- The mean excess function (MEF) to discriminate between the classes D and E; and
- Two statistics: Hill's ratio and modified Jackson statistic, for confirmatory analysis of the conclusions suggested by the previous two methods.
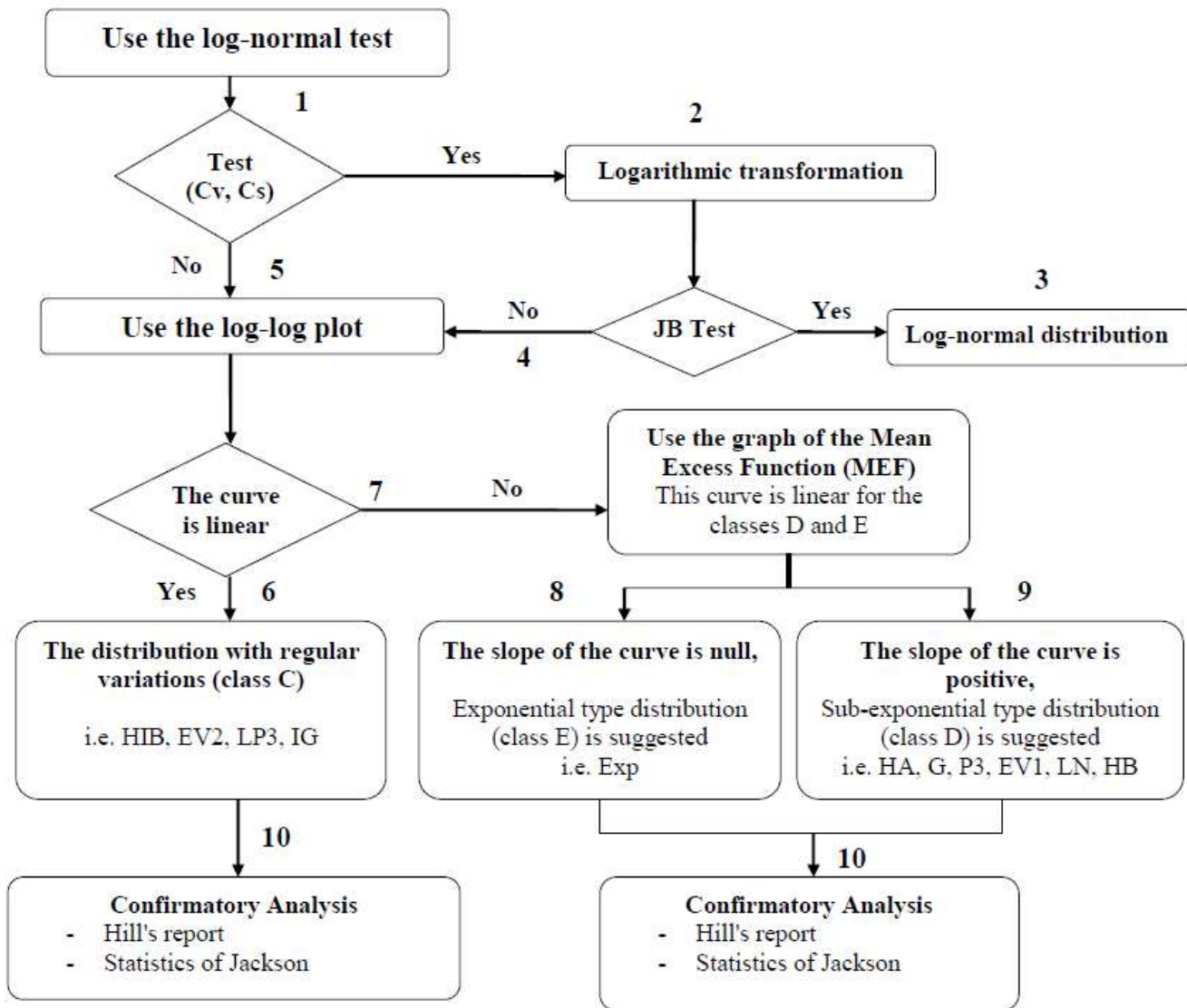
Figure 2: Diagram for class discrimination used in the DSS

More theoretical details of this classification and the criteria are available in El Adlouni et al. (2008) and Martel et al. (2011). These articles are available as attachment in the HYFRAN-PLUS setup.

2. Log-normality test

2.1. The (Cv,Cs) diagram

To test the log-normality, tests of normality are applied to the logarithmic transformation Y of the initial data (Y = Log (X), where X is the original variable). Recent work has been done by Martel, El Adlouni and Bobée (2011), to define a procedure based on one or more tests to discriminate between the Log-normale distribution and the classes C and D. Five tests (Anderson-Darling (AD), Shapiro-Wilk (SW), Lilliefors (Lf), Jarque-Bera (JB) and Filliben (FB)) were compared by simulation to examine their powers. Note that the five tests used to test normality, were used on the series transformed by the logarithmic function. Martel, El Adlouni and Bobée (2011) [Article available as an attached file when installing HYFRAN-PLUS] showed that when the series has examined some characteristics (represented by coefficients of variation Cv and skewness Cs) the test of Jarque-Bera (JB) has a satisfactory power to test the log-normality, when the alternative is a distribution of the class C or D. Indeed, if the coefficients of skewness and variation of the studied series are in the area corresponding to the Inverse Halphen type B (HIB) of the diagram (Cv,Cs) (Figure 3), the power of the JB test is very satisfactory and its use is recommended in the DSS.
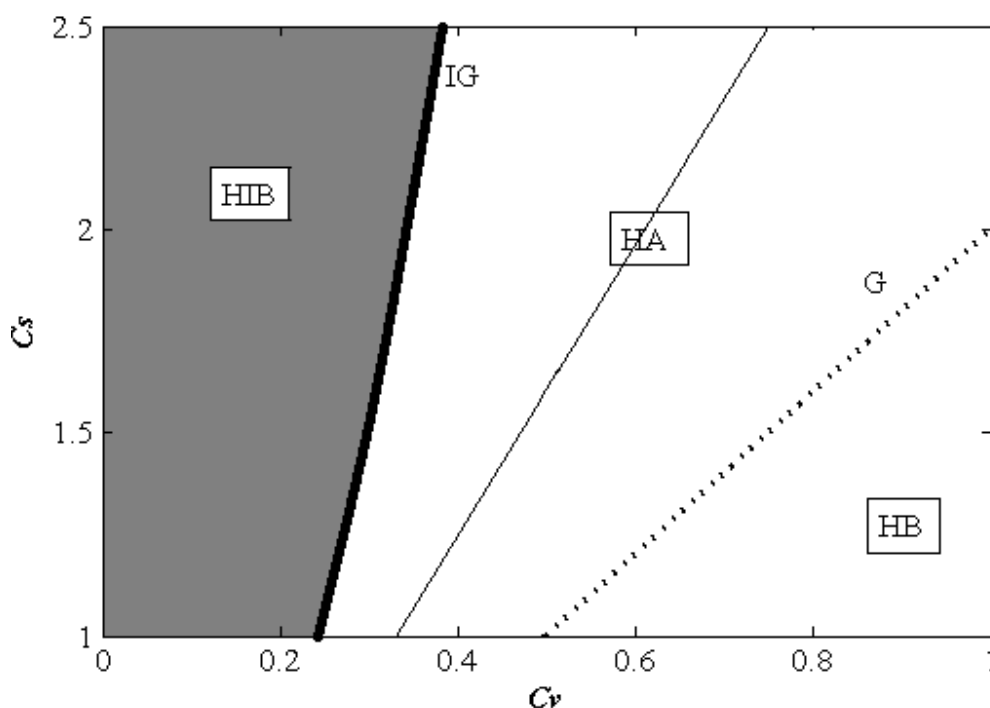

Figure 3 : The HIB region in the (Cv,Cs) diagram

We therefore propose as a first step in the DSS, the use of the (Cv, Cs)-diagram and the JB test. The parts of this step are summarized by the following algorithm:

1. Compute the Cv and Cs of the sample [Step 1];
2. If the point (Cv,Cs) belongs to the Halphen Inverse B (HIB) [Step 2] (Figure 3): Apply the JB test on the transformed sample (using the logarithmic function) to test normality.
3. If the Log-normality is accepted at significance level of 5%, the Log-normal distribution is recommended to fit the dataset [Step 3].
4. If the log-normality hypothesis (normality of the transformed data) is rejected, at significance level 5%, use the Lo-log diagram (Section 3) [Step 4].
5. If the sample does not belong to the zone HIB: we do not consider the LN distribution to fit the data and suggest the of the Log-log diagram (Section 3) [Step 5].

2.2. The Jarque-Bera Test [JB; Jarque et Bera, 1980]

The JB test (Jarque and Bera, 1980) uses functions of the 3rd and 4th moments of the sample, which correspond to the coefficients of skewness (Cs) and kurtosis (Ck). In the case of a normal distribution N(.;.), these coefficients are respectively equal to 0 and 3. JB test combines these two factors into a single statistic:

$$JB = N\left\{\frac{1}{6}(Cs)^2 + \frac{1}{24}(Ck-3)^2\right\}$$

Where N is the sample size, with the condition $N > 7$. The JB statistic follows a chi-square distribution χ_2^2, with degrees of freedom $\nu = 2$. It then compares the value the calculated JB statistic from the sample to the critical value at significance level of 5%. The decision rule is:

- If $JB < \chi^2_{2,0.95}$, the H0 hypothesis is accepted (the transformed series is normal) and therefore we accept the log-normality hypothesis[Step 3];

- If $JB > \chi^2_{2,0.95}$ the H0 hypothesis is rejected (the transformed series is not normal) and we therefore reject the hypothesis of log-normality [Step 4].

It should be noted that this is a test and is based on asymptotic result, therefore, generally not very effective for small samples ($N < 30$). We decided, as part of DSS, use it for samples, such as coefficients Cv, Cs belong to the region HIB. For this region of the (Cv,Cs) diagram, the JB test has an acceptable

power even for small sample size (Martel El Adlouni and Bobée, 2011).

3. Log-Log plot

The log-log plot is based on the fact that the survival function $\bar{F}(u) = P(X > u)$, is given by $\bar{F}(u) = P(X > u) = e^{-u/\theta}$ for exponential tail with mean θ, and for regularly varying distribution with tail index α, $\bar{F}$ is equivalent to (for large quantile) :

$$\bar{F}(u) = P(X > u) \approx C \int_u^\infty \frac{1}{x^\alpha} dx = C \left[\frac{x^{-\alpha+1}}{1-\alpha} \right]_u^\infty = C_1 u^{-\alpha+1}$$ (with $\alpha > 1$, which is equivalent to finite mean).

Therefore, taking the logarithm we have regularly varying distributions $\log[P(X > u)] \approx \log C_\alpha - (\alpha-1)\log(u)$. This suggests that, for the log-log plot, the tail probability is represented by a straight line for power-law (or regularly varying distributions, class C) but not for the other sub-exponential or exponential distributions (class D or E).

As illustrated in Figure 4, the curve represented in the Log-Log plot corresponds to a straight line for the distributions of the class C i.e. Fréchet (EV2), Halphen type IB (HIB), Log-Pearson type 3 (LP3) and Inverse Gamma (IG), but not for sub-exponential or exponential type tails (class D or E). When the diagram is not linear we suggest the use of the Mean Excess Function (MEF) to discriminate between the classes D and E.
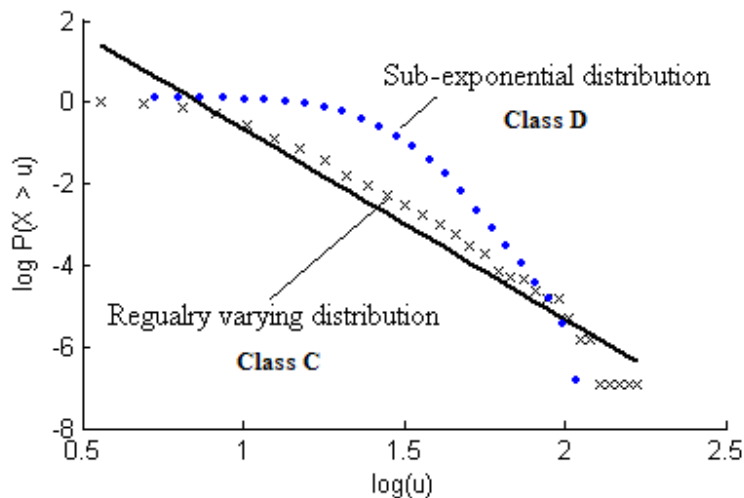

Figure 4: Illustration of the Log-Log plot to characterize the regularly varying distributions

To check the linearity of the curve in the log-log diagram, a test on the associated correlation coefficient is considered. Simulation studies allow the determination of critical values corresponding to significance levels of 5 % and 1 %, to test the HYPOTHESIS H0: THE DATA FOLLOW A DISTRIBUTION OF THE CLASS C (i.e. THE CURVE IS LINEAR). These critical values are calculated according to the size N of the sample ($30 \leq N \leq 200$). ***Note that the decisions given by the DSS are based, by default, on the significance level 5 %.***

If the hypothesis H0 is rejected, at the significance level 5 %, we suggest the use of the mean excess function plot (MEF). ***However the critical values at the significance level 1 % are given for more flexibility and to allow the user to make another decision than that based on the significance level 5 %.***

Indeed, if the observed correlation coefficient (ro) is greater than critical value (rc) at the significance level 5 %, then we conclude that it is not significantly different from 1 at the significance level 5 % and the hypothesis H0 of linearity is accepted at this level (Figure 4). In this case, the most adequate choice corresponds to the class C of regularly varying distributions (power-law type) : Halphen type IB (HIB), Fréchet (EV2), Log-Pearson type 3 (LP3), Inverse Gamma (IG).
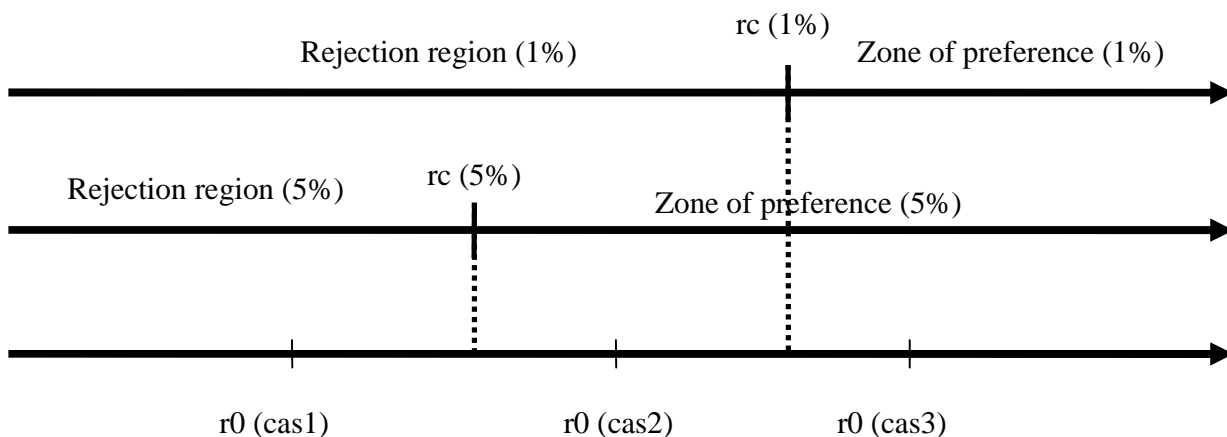

Figure 5: Illustration of the one-sided test of the hypothesis H0.

Figure 5 shows, in general, the decision rule for a one-sided test for to two significance levels 1% and 5%. The critical values corresponding to each significance level are, respectively, rc(1%) and rc(5%). These two critical values are obtained by Monte Carlo simulations generated from regularly varying distributions. For a given dataset, we calculate the correlation coefficient r0. To illustrate the use of this

test, three cases are considered such as the correlation coefficients verify: r0(cas1) < rc(5%) < r0(cas2) < rc(1%) < r0(cas3). The hypothesis H0 (case1) is rejected for the significance levels 1% and 5% . Indeed, r0(cas1) < rc(5%) and r0(cas1) < rc(1%). In this case the distribution is not regularly varying (the curve is not linear). For case2, the hypothesis H0 is rejected at the significance level 1%, but it is accepted at the significance level of 5%. Indeed, r0(cas2) > rc(5%) and r0(cas2) < rc(1%). For this case, the hypothesis H0 is accepted by the SAD and the use of regularly varying distribution is suggested (based on the significance level 5%). However, the critical value at the significance level of 1% is presented to give more flexibility to the user. The case 3, corresponds to the case where r0 is higher than the two critical values (r0(cas3) > rc(5%) and r0(cas3) > rc(1%)). In this case, and for the two significance levels, the hypothesis H0 is accepted and the suggested distribution belong to the class C of regularly varying distributions.

4. The Mean Excess Function Diagram (MEF)

The mean excess function method is based on the function $e(u) = E[X - u \,|\, X > u]$. This function is constant for exponential tail distributions ($e(u) = \theta$). However, in the case of regularly varying distribution with tail index $\alpha \,(\alpha > 2)$: $e(u) = \dfrac{u}{(\alpha - 2)}$. The Mean Excess Function (MEF) allows discriminating between the class D (sub-exponential distributions) and the class E (Exponential distribution). Indeed, the curve presented in the MEF diagram is linear for high observed values for distributions of both classes D and E. If in addition the slope of this curve is (Figure 6):

- Equal to zero, the most adequate distribution belongs to the class E (Exponential law);
- Strictly positive, the most adequate distribution belongs to the class D of sub-exponential distributions: Halphen type A (HA), Gumbel (EV1), Halphen type B (HB), Pearson type III (PIII), Gamma (G).

Note that, in the SAD, this method should be used after the log-log plot method. Indeed, if the assumption H0 of the log-log method is rejected (distribution does not belongs to the class C of regularly varying distributions) FME method allows testing whether the distribution is exponential or not.
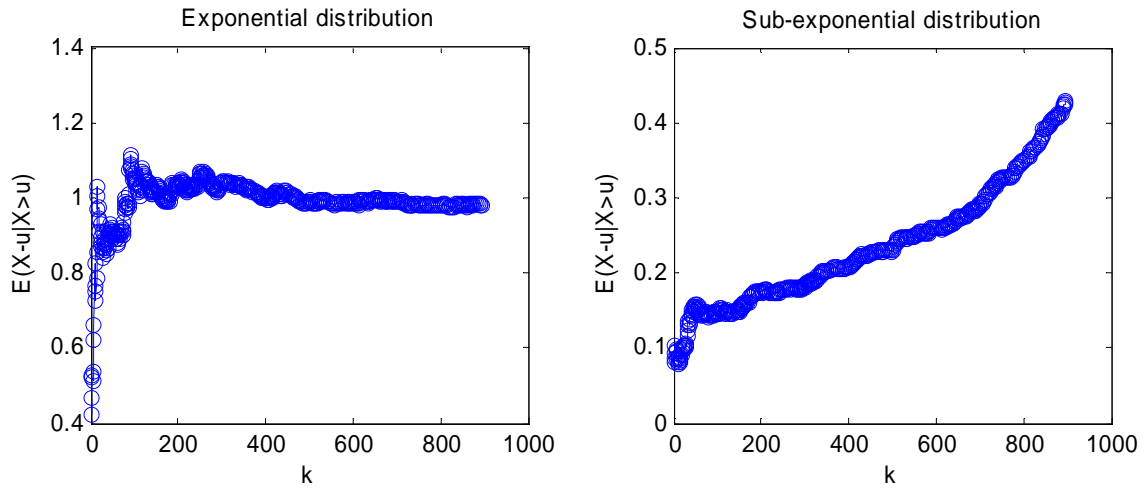
Figure 6: Mean excess function for exponential and sub-exponential distributions.

The use of this diagram in the DSS is based on the slope of the MEF curve for the observations that exceed the median (50 % of the highest observed value of the sample).

Simulation studies allow the determination of critical values of the slope corresponding to significance levels of 5 % and 1 %, to test the HYPOTHESIS H0: THE DATA FOLLOW A DISTRIBUTION OF THE CLASS E (i.e. THE SLOPE OF THE MEF IS EQUAL TO ZERO). These critical values are calculated according to the size N of the sample ($30 \leq N \leq 200$). ***Note that the decisions given by the DSS are based, by default, on the significance level 5 %.***

When the hypothesis H0 is accepted we suggest the use of the Exponential distribution (class E). However, when it is rejected at the significance level 5 %, we suggest the use of a distribution of the class D (HA, EV1, HB, PIII, G). ***Note that the critical values at the significance level 1 % are given for more flexibility and to allow the user to make possibly another decision than that suggested for the significance level of 5%.*** The zone of preference or rejection of the null hypothesis for the significance levels 1% and 5%, can be represented in a similar manner as for the Log-log method (Figure 5).

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5. Hill's ratio plot [for the theoretical details cf. El Adlouni et al. 2008]

The Hill ratio is defined by

$$a_n(x_n) = \frac{\sum_{i=1}^{n} I(X_i > x_n)}{\sum_{i=1}^{n} \log(X_i / x_n) I(X_i > x_n)}$$

where $I(X_i > x_n) = \begin{cases} 1 & \text{if} & X_i > x_n \\ 0 & \text{if} & X_i < x_n \end{cases}$.

This method is based on the fact that a_n is a consistent estimator of α if the tail is regularly varying (Class C) with tail index α (Hill, 1975). In the expression of the Hill ratio, x_n is chosen to be large such that $P(X > x_n) \to 0$ and $nP(X > x_n) \to \infty$, and I is the indicator function. The standard Hill estimator, of the tail index, corresponds to the particular case where the observations are ordered $X_{(1)} \le \ldots \le X_{(n)}$ and $x_n = X_{(k_n+1)}$, where k_n is an integer which tends to infinity as n tends to infinity.

In practice, one plots $a_n(x_n)$ as a function of x_n and looks for some stable region from which $a_n(x_n)$ can be considered as an estimator of α. Figure 7, presents the Hill ratio plot for a sample generated from the regularly varying (a) and Exponential (b) distributions.
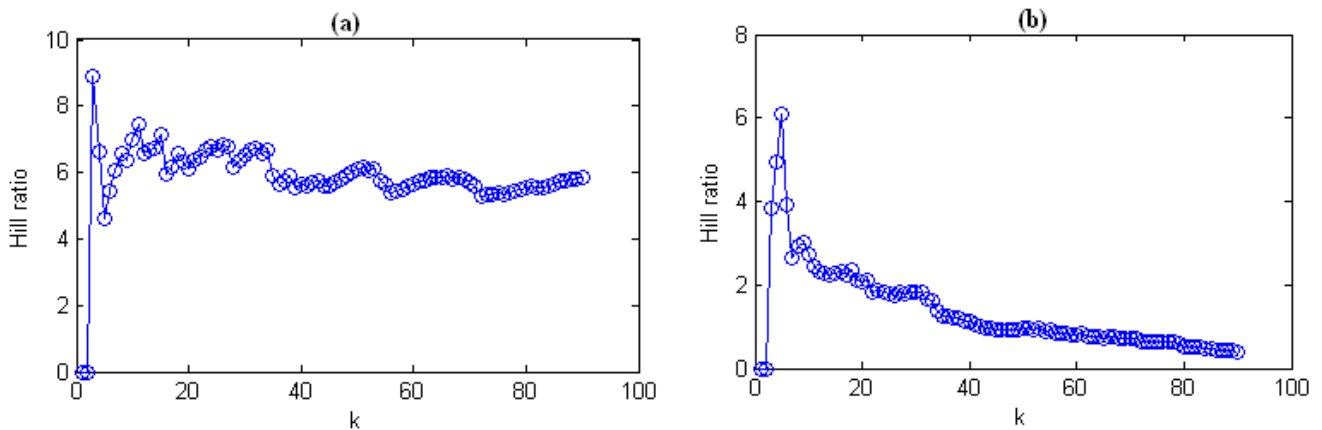

Figure 7: Generalized Hill ratio plot for (a) regularly-varying and (b) sub-exponential distributions.

This statistics is used in the DSS to confirm the suggested choice given by the first two diagrams (the distribution belongs to the class C, D or E).

- If the curve converges to a non-null constant value, the most adequate distribution belongs to the class C (regularly varying distribution). We suggest then the use of a distribution of the class C: Fréchet (EV2), Halphen type B Inverse (HIB), Log-Pearson type 3 (LP3), Inverse Gamma (IG).

- If the curve **decreases to zero**, the distribution belong to the Sub-exponential class (class D: Halphen type A, Gamma, Pearson type III, Halphen type B, Gumbel); and the Exponential class (class E: Exponential distribution).

Note that (cf. section 3) to discriminate between the classes D and E, we suggest the use of the MEF method.

6. Jackson Statistic [for the theoretical details cf. El Adlouni et al. 2008]

This method is presented by Beirlant et al. (2006) and is based on the Jackson statistic. It allows to test whether the sample is consistent with Pareto type distributions. Note that the distributions of the class C (regularly varying distribution) have asymptotically the same behaviour as that of the Generalized Pareto distribution. Originally the Jackson statistic (Jackson, 1967) was proposed as a goodness-of-fit statistic for testing exponential behaviour, and given the link between the Exponential and the Pareto distribution (if X has a Pareto distribution the logarithmic transformation $Y = \log(X)$ is exponentially distributed) this statistic is used to assess Pareto-type behaviour. The Jackson statistic is further modified by taking into account the second-order tail behaviour of a Pareto-type model. Beirlant et al. (2006) give the limiting distribution of this statistic with corrected bias version for finite size samples. The modified Jackson statistic converges to 2 for power tail type distribution and has an irregular behaviour for sub-exponential or exponential distributions (Figure 8). In the DSS, the Jackson statistic is used to characterize distributions of the class C. Indeed, regularly varying distributions (class C) has asymptotically a power tail.
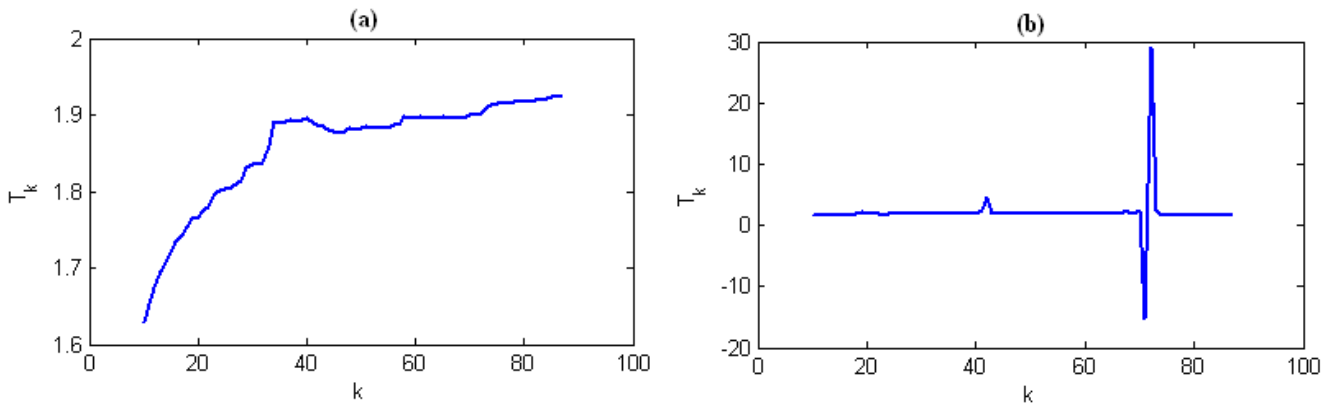
Figure 8: Modified Jackson statistic for (a) regularly varying and (b) sub-exponential distributions.

In the DSS this method is considered as a confirmatory method for suggested decision based on the Log-Log and the MEF. So:

- If the curve converges clearly and regularly to 2 (figure 8-a), the studied distribution belongs to the class C (regularly varying distribution). We suggest then, the use of: Fréchet (EV2), Halphen type IB (HIB), Log-Pearson type 3 (LP3), Inverse Gamma (IG);

- If the curve presents some irregularities and do not converge to 2 (figure 8-b), than we suggest the sub-exponential class (class D: Halphen type A, Gamma, Pearson type III, Halphen type B, Gumbel); or exponential (class E: Exponential distribution).

Note that (cf. section 3) to discriminate between the classes D and E, we suggest the use of the MEF method.

Remarque:

Even if the modified Jackson statistic was developed to test Pareto type behaviour, it is used in the DSS to check if the of the studied distribution has similar tail as regularly varying distribution (class C). In deed, distributions of the class C have asymptotically Pareto type tail. In practice, the Generalized Pareto distribution (GPD) is used in the Peaks-over-threshold model (POT). However, the GPD is available in HYFRAN and can be used to fit any data sets that are independent, homogenous and stationary.

Reference:

Beirlant, J., de Wet, T., Goegebeur, Y., (2006). A goodness-of-fit statistic for Pareto-type behaviour. Journal of Computational and Applied Mathematics, 186, 99-116.

El Adlouni, S., Bobée, B. et Ouarda, T. B.M.J (2008). On the tails of extreme event distributions in Hydrology. Accepted in Journal of Hydrology.

Jackson, O.A.Y., (1967). An analysis of departures from the exponential distribution. Journal of the Royal Statistical Society B, 29, 540-549.

Martel, B., S. El Adlouni et B. Bobée. Comparison of the power of Log-Normality tests with different right tail alternative distributions. J. Hyd. Eng. ASCE, 2011, Soumis.

Jarque, C. M., and Bera, A. K. (1987). "A Test for Normality of Observations and Regression Residuals." International Statistical Review / Revue Internationale de Statistique, 55(2), 163-172.